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If you **fully understand the toy model**, you should find the more realistic application straightforward.

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As an introduction to statistical estimation, we consider the simple problem, that we call the **two temperatures problem**:

Given two independent observations T_1 and T_2 , determine the best estimate of the true temperature T_t .

Simple (toy) Example

Let the two observations of temperature be

$$\left. \begin{aligned} T_1 &= T_t + \varepsilon_1 \\ T_2 &= T_t + \varepsilon_2 \end{aligned} \right\}$$

[For example, we might have two *iffy* thermometers].

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The observations have **errors** ε_i , which we don't know.

Let $E(\)$ represent the **expected value**, i.e., the average of many similar measurements.

We assume that the measurements T_1 and T_2 are **unbiased**:

$$E(T_1 - T_t) = 0, \quad E(T_2 - T_t) = 0$$

or equivalently,

$$E(\varepsilon_1) = E(\varepsilon_2) = 0$$

We also assume that we know the **variances** of the observational errors:

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The above equations represent the statistical information that we need about the actual observations.

We estimate T_t as a **linear combination** of the observations:

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T_a will be the **best estimate** of T_t if the coefficients are chosen to **minimize the mean squared error** of T_a :

$$\sigma_a^2 = E[(T_a - T_t)^2] = E\{[a_1(T_1 - T_t) + a_2(T_2 - T_t)]^2\}$$

subject to the **constraint** $a_1 + a_2 = 1$.

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This may be written

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Expanding this expression for σ_a^2 , we get

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Naïve solution: $\partial\sigma_a^2/\partial a_1 = 2a_1\sigma_1^2 = 0$, so $a_1 = 0$.

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Thus, we have expressions for the **weights** a_1 and a_2 in terms of the **variances** (which are assumed to be known).

We define the **precision** to be the inverse of the variance. It is a measure of the accuracy of the observations.

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This can be written in the alternative form:

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Thus, *if the coefficients are optimal, the **precision of the analysis** is the **sum of the precisions** of the measurements.*

Variational approach

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The **cost function** is defined as the sum of the squares of the distances of T to the two observations, weighted by their observational error precisions:

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Exercise: Prove that $\partial J / \partial T = 0$ gives the same value for T_a as the least squares method.

The **control variable** for the minimization of J (i.e., the variable with respect to which we are minimizing the cost function) is the **temperature**.

For the least squares method, the control variables were **the weights**.

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The **equivalence** between the minimization of the analysis error variance and the variational cost function approach is important.

This equivalence also holds true for multidimensional problems, in which case we use the **covariance matrix** rather than the scalar variance.

It indicates that OI and 3D-Var are solving the same problem by different means.

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Example: Suppose $T_1 = 2$ $\sigma_1 = 2$ $T_2 = 0$ $\sigma_2 = 1$.

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$$a_1 = \frac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2} = \frac{1}{5} \qquad a_2 = \frac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2} = \frac{4}{5}$$

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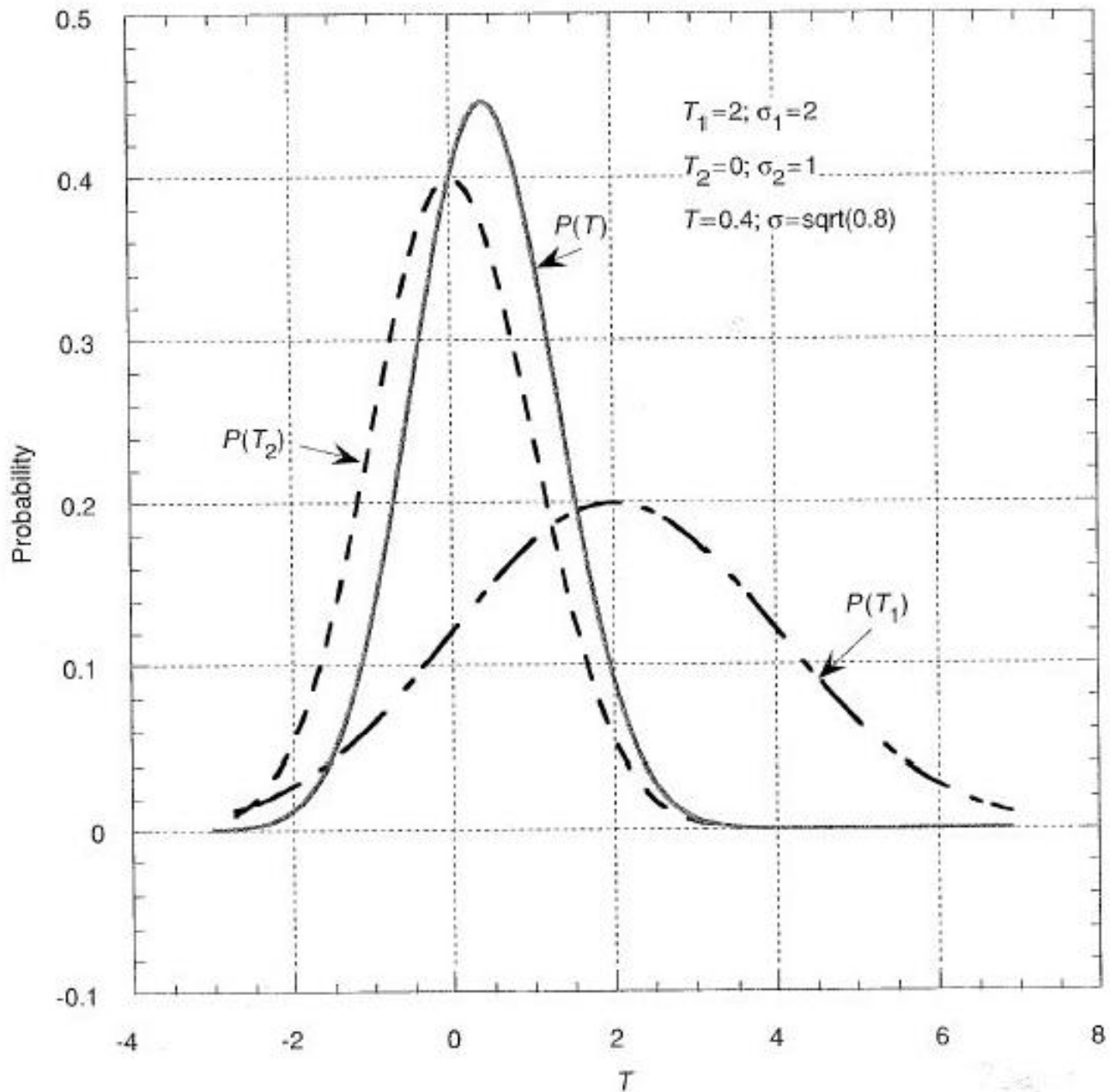
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This solution is illustrated in the next figure.



The probability distribution for a simple case.

The analysis has a pdf with a maximum closer to T_2 , and a smaller standard deviation than either observation.

Conclusion of the foregoing.

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Assume that one of the two temperatures, say $T_1 = T_b$, is not an observation, but a **background value**, such as a forecast.

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Assume that the other value is an **observation**, $T_2 = T_o$.

We can write the analysis as

$$T_a = T_b + W(T_o - T_b)$$

where $W = a_2$ can be expressed in terms of the variances.

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the quantity $(T_o - T_b)$ is called the *observational innovation*, i.e., the **new information** brought by the observation.

It is also known as the **observational increment** (with respect to the background).

The analysis error variance is, as before, given by

$$\frac{1}{\sigma_a^2} = \frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \quad \text{or} \quad \sigma_a^2 = \frac{\sigma_b^2 \sigma_o^2}{\sigma_b^2 + \sigma_o^2}$$

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We have shown that the simple **two-temperatures** problem serves as a **paradigm** for the problem of objective analysis of the atmospheric state.

Collection of Main Equations

We gather the principal equations here:

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$$T_a = T_b + W(T_o - T_b)$$

$$W = \frac{\sigma_b^2}{\sigma_b^2 + \sigma_o^2}$$

$$\sigma_a^2 = \frac{\sigma_b^2 \sigma_o^2}{\sigma_b^2 + \sigma_o^2} = W \sigma_o^2$$

$$\sigma_a^2 = (1 - W) \sigma_b^2$$

These four equations have been derived for the simplest **scalar case** ...

... but they are important for the problem of data assimilation because they have exactly the same form as **more general equations**:

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The least squares sequential estimation method is used for real multidimensional problems (OI, interpolation, 3D-Var and even Kalman filtering).

Therefore we will interpret these four equations in detail.

The first equation

$$T_a = T_b + W(T_o - T_b)$$

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This says:

The analysis is obtained by adding to the background value, or first guess, the innovation (the difference between the observation and first guess), weighted by the optimal weight.

The second equation

$$W = \frac{\sigma_b^2}{\sigma_b^2 + \sigma_o^2}$$

This says:

The optimal weight is the background error variance multiplied by the inverse of the total error variance (the sum of the background and the observation error variances).

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Look at the limits: $\overset{\star}{\sigma_o^2} = 0$; $\overset{\star}{\sigma_b^2} = 0$.

The third equation

The variance of the analysis is

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This says:

The precision of the analysis (inverse of the analysis error variance) is the sum of the precisions of the background and the observation.

The fourth equation

$$\sigma_a^2 = (1 - W)\sigma_b^2$$

This says:

The error variance of the analysis is the error variance of the background, reduced by a factor equal to one minus the optimal weight.

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Note that there is one essential *tuning* parameter in OI:

It is the ratio of the observational variance to the background error variance:

$$\left(\frac{\sigma_o}{\sigma_b}\right)^2$$

Application to Analysis

If the background is a forecast, we can use the four equations to create a simple sequential **analysis cycle**.

The observation is used once at the time it appears and then discarded.

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Assume that we have completed the analysis at time t_i (e.g., at 06 UTC), and we want to proceed to the next cycle (time t_{i+1} , or 12 UTC).

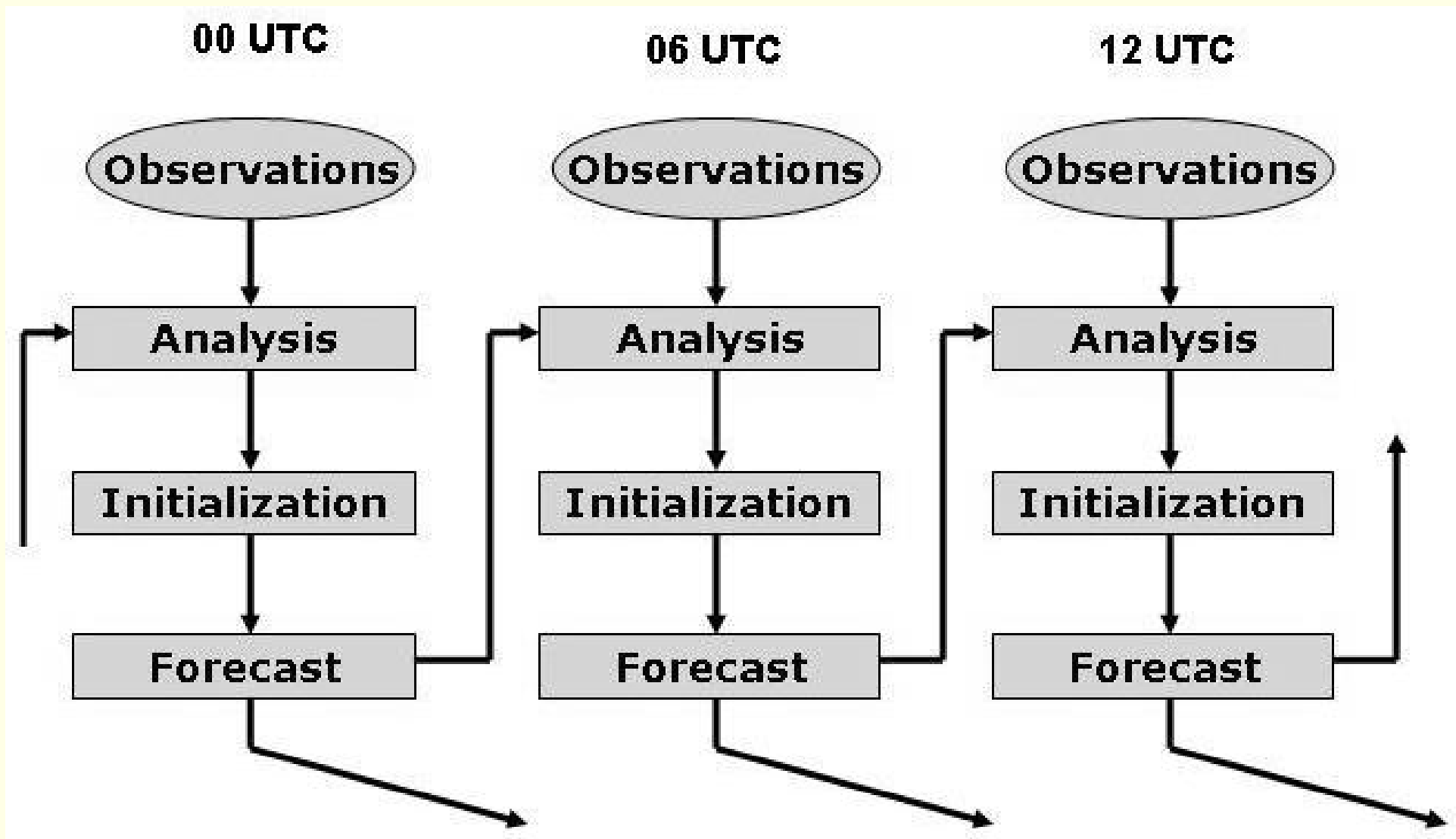
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The analysis cycle has two phases, a **forecast** phase to update the background T_b and its error variance σ_b^2 , and an **analysis** phase, to update the analysis T_a and its error variance σ_a^2 .



Typical 6-hour analysis cycle.

Forecast Phase

In the **forecast phase of the analysis cycle**, the background is first obtained through a forecast:

$$T_b(t_{i+1}) = M [T_a(t_i)]$$

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This allows the new weight $W(t_{i+1})$ to be estimated using

$$W = \frac{\sigma_b^2}{\sigma_b^2 + \sigma_o^2}$$

Analysis Phase

In the **analysis phase of the cycle** we get the new observation $T_o(t_{i+1})$, and we derive the new analysis $T_a(t_{i+1})$ using

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After the analysis, the cycle for time t_{i+1} is completed, and we can proceed to the next cycle.

Reading Assignment

Study the **Remarks** in Kalnay, §5.3.1